

Supplementary Material for “Online Estimation for Functional Data”

In this Supplementary Material, we first derive the asymptotic distributions of the proposed mean and covariance estimates, i.e., Theorem 1 and 2, in Section S.1 and S.2, respectively. The convergence of bandwidth selection and lower bound of relative efficiencies are shown in Section S.3 and S.4.

S.1. PROOF OF THEOREM 1

Proof. This proof is in analogy to the proof of Theorem 3.1 in Zhang and Wang (2016). Denote the current time by K and recall $\{\tilde{\eta}_{\mu,k}^{(K)}; k = 1, \dots, K\}$ is the pseudo-bandwidth chain at time K . We omit the superscript “ (K) ” of $\tilde{\eta}_{\mu,k}^{(K)}$, $\tilde{h}_{\mu}^{(K)}$ in (18), $\rho_{\mu,j}^{(K)}$ in (14) and $\tilde{\mu}^{(K)}$ in (10) in the proof when no confusion exists. The online estimate can be written as

$$\tilde{\mu}(t) = \frac{R_0 Q_2 - R_1 Q_1}{Q_0 Q_2 - Q_1^2}$$

where for $r = 0, 1, 2$,

$$Q_r = \frac{1}{S_{K,1}} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{j=1}^{m_{ki}} W_{\tilde{\eta}_{\mu,k}}(T_{kij} - t) \left(\frac{T_{kij} - t}{\tilde{\eta}_{\mu,k}} \right)^r = \sum_k \omega_k Q_{r,k},$$

$$R_r = \frac{1}{S_{K,1}} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{j=1}^{m_{ki}} W_{\tilde{\eta}_{\mu,k}}(T_{kij} - t) \left(\frac{T_{kij} - t}{\tilde{\eta}_{\mu,k}} \right)^r Y_{kij} = \sum_k \omega_k R_{r,k}.$$

We further denote for $r = 0, 1$,

$$\tilde{G}_r = \frac{1}{S_{K,1}} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{j=1}^{m_{ki}} W_{\tilde{\eta}_{\mu,k}}(T_{kij} - t) T_{kij}^r = \sum_k \omega_k \tilde{G}_{r,k},$$

and define

$$\tilde{\mu}'(t) = \frac{1}{\tilde{\eta}_{\mu,k}} \frac{R_1 - R_0 Q_1 / Q_0}{Q_2 - Q_1^2 / Q_0}.$$

Note that

$$Q_1 = \sum_k \omega_k Q_{1,k} = \sum_k \omega_k \frac{\tilde{G}_{1,k} - t \tilde{G}_{0,k}}{\tilde{\eta}_{\mu,k}},$$

one can derive

$$\tilde{\mu}(t) = \frac{R_0}{Q_0} - \sum_k \omega_k \frac{\tilde{G}_{1,k}}{Q_0} \tilde{\mu}'_k(t) + t \sum_k \omega_k \frac{\tilde{G}_{0,k}}{Q_0} \tilde{\mu}'_k(t).$$

We also define

$$\bar{\mu}(t) = \frac{R_0}{Q_0} - \sum_k \omega_k \frac{\tilde{G}_{1,k}}{Q_0} \mu'(t) + t \sum_k \omega_k \frac{\tilde{G}_{0,k}}{Q_0} \mu'(t), \quad (\text{S.1})$$

where $\mu'(\cdot)$ is the first derivative of $\mu(\cdot)$. Then

$$\tilde{\mu}(t) - \bar{\mu}(t) = - \sum_k \omega_k \frac{Q_{1,k}}{Q_0} \cdot \frac{Q_0(R_1 - \mu Q_1 - \tilde{\eta}_{\mu,k} \mu' Q_2) - Q_1(R_0 - \mu Q_0 - \tilde{\eta}_{\mu,k} \mu' Q_1)}{Q_2 Q_0 - Q_1^2}. \quad (\text{S.2})$$

It is straightforward to show that both Q_0 and $Q_0 Q_2 - Q_1^2$ are positive and bounded away from 0 with probability tending to one,

$$Q_1 = O_p \left\{ \rho_{\mu,1} + \left(\frac{\rho_{\mu,-1}}{S_{K,1}} \right)^{\frac{1}{2}} \right\},$$

$$(R_1 - \mu Q_1 - \tilde{\eta}_{\mu,k} \mu' Q_2), (R_0 - \mu Q_0 - \tilde{\eta}_{\mu,k} \mu' Q_1) = O_p \left(\rho_{\mu,2} + \rho_{\mu,1} \tilde{\eta}_{\mu,k} + \left\{ \frac{\rho_{\mu,-1}}{S_{K,1}} \right\}^{\frac{1}{2}} + \left\{ \frac{S_{K,2}}{S_{K,1}^2} \right\}^{\frac{1}{2}} \right).$$

Then by the Cramér-Wold device and Lyapunov condition, we can achieve the asymptotic joint normality of $(R_0 - \text{E}R_0, \tilde{G}_1 - \text{E}\tilde{G}_1, \tilde{G}_0 - \text{E}\tilde{G}_0)$, where the rate of convergence is $\min[(\rho_{\mu,-1}/S_{K,1})^{1/2}, \{Q_2/S_{K,1}^2\}^{1/2}]$.

Explicitly for $r, r' = 0, 1$,

$$\text{E}(\tilde{G}_r) = t^r f(t) + \frac{1}{2} \alpha(W) \{2r f'(t) + t^r f''(t)\} \rho_{\mu,2} + o_p(\rho_{\mu,2}),$$

$$\text{E}(R_0) = \mu(t) f(t) + \frac{1}{2} \alpha(W) \{\mu(t) f''(t) + 2\mu'(t) f'(t) + \mu''(t) f(t)\} \rho_{\mu,2} + o_p(\rho_{\mu,2}),$$

$$\text{Cov}(\tilde{G}_r, \tilde{G}_{r'}) = \left(\frac{\rho_{\mu,-1}}{S_{K,1}} \right) R(W) t^{r+r'} f(t) + o_p \left(\frac{\rho_{\mu,-1}}{S_{K,1}} \right),$$

$$\text{Var}(R_0) = \left(\frac{\rho_{\mu,-1}}{S_{K,1}} \right) R(W) (\mu^2(t) + \gamma(t, t) + \sigma^2) f(t) + \left(\frac{S_{K,2}}{S_{K,1}^2} \right) \gamma(t, t) f(t)^2 + o_p \left(\frac{\rho_{\mu,-1}}{S_{K,1}} + \frac{S_{K,2}}{S_{K,1}^2} \right),$$

$$\text{Cov}(R_0, \tilde{G}_r) = \left(\frac{\rho_{\mu,-1}}{S_{K,1}} \right) R(W) t^r \mu(t) f(t) + o_p \left(\frac{\rho_{\mu,-1}}{S_{K,1}} \right).$$

From the delta method, $\bar{\mu}(t)$ follows the asymptotic normality that

$$\Gamma_\mu^{-1/2} \left\{ \bar{\mu}(t) - \mu(t) - \frac{1}{2} \alpha(W) \mu^{(2)}(t) \rho_{\mu,2} + o_p(\rho_{\mu,2}) \right\} \xrightarrow{d} N(0, 1),$$

where

$$\Gamma_\mu = R(W) \frac{\gamma(t, t) + \sigma^2}{S_{K,1} f(t)} \rho_{\mu,-1} + \gamma(t, t) \frac{S_{K,2}}{S_{K,1}^2}.$$

When $\bar{m}_{\mu,K}/(N_K)^{1/4} \rightarrow 0$ and $\tilde{h}_\mu \asymp S_{K,1}^{-1/5}$, the first term in Γ_μ above dominates. When $\bar{m}_{\mu,K}/(N_K)^{1/4} \rightarrow C$ and $\tilde{h}_\mu \asymp S_{K,1}^{-1/5}$, where $0 < C < \infty$, the two terms of Γ_μ are of the same

order as \tilde{h}_μ^4 . When $\bar{m}_{\mu,K}/(N_K)^{1/4} \rightarrow \infty$ and $\tilde{h}_\mu \asymp S_{K,1}^{-1/5}$, the second term of Γ_μ dominates and the bias becomes negligible. Noting (S.2), the proof of Theorem 1 is completed. \blacksquare

S.2. PROOF OF THEOREM 2 AND LEMMA 1

We omit the superscript “(K)” of $\tilde{\eta}_{\gamma,k}^{(K)}$, $\tilde{h}_\gamma^{(K)}$, $\rho_{\gamma,j}^{(K)}$ and $\tilde{\gamma}^{(K)}$ in this section when no confusion exists. The proof of Lemma 1 is similar to the proof of Theorem 2, with $\rho_{\gamma,j}$, $j \in \mathbb{Z}$ replaced by \hat{h}_γ . Before delineating the proof of Theorem 2, we state the following lemma which is a preparation for the proof of Theorem 2.

Lemma S. 1. *Let*

$$Q_{0R_1} = \frac{1}{S_{K,1}} \sum_{k'=1}^K \sum_{i'=1}^{n_{k'}} \sum_{j'=1}^{m_{k'i'}} W_{\tilde{\eta}_{\mu,k'}}(T_{k'i'j'} - T_{kil}),$$

$$\Delta \tilde{R}_1 = -S_{K,1}^{-1} S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{j \neq l}^{n_k} \sum_{k'=1}^K \sum_{i'=1}^{n_{k'}} \sum_{j'=1}^{m_{k'i'}} W_{\tilde{\eta}_{\gamma,k'}}(T_{kij} - s) W_{\tilde{\eta}_{\gamma,k'}}(T_{kil} - t) W_{\tilde{\eta}_{\mu,k'}}(T_{k'i'j'} - T_{kil}),$$

$$\cdot (Y_{kij} - \mu(T_{kij})) \left\{ Y_{k'i'j'} - \mu(T_{k'i'j'}) - \frac{1}{2} \mu''(T_{kil})(T_{k'i'j'} - T_{kil})^2 \right\},$$

then

$$\begin{aligned} \mathbb{E}(\Delta \tilde{R}_1 / Q_{0R_1}) &= O_p(S_{K,1}^{-1} S_{K,2}^{-1} S_{K,3}), \\ \text{Var}(\Delta \tilde{R}_1 / Q_{0R_1}) &= O_p \left(S_{K,1}^{-2} S_{K,2}^{-2} \left\{ S_{K,3}^2 + \sum_k s_{k,3} \tilde{\eta}_{\mu,k}^{-1} \tilde{\eta}_{\gamma,k}^{-2} + \sum_k s_{k,4} \tilde{\eta}_{\mu,k}^{-1} \tilde{\eta}_{\gamma,k}^{-1} \right. \right. \\ &\quad \left. \left. + \sum_k s_{k,5} \tilde{\eta}_{\gamma,k}^{-1} \right\} + S_{K,2}^{-2} \sum_k s_{k,4} \tilde{\eta}_{\gamma,k}^4 \right). \end{aligned}$$

Proof. By easy calculation,

$$\begin{aligned} \mathbb{E} \Delta \tilde{R}_1 &= -S_{K,1}^{-1} S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{j \neq l}^{n_k} \sum_{k'=1}^K \sum_{i'=1}^{n_{k'}} \sum_{j'=1}^{m_{k'i'}} \mathbb{E} \left\{ W_{\tilde{\eta}_{\gamma,k'}}(T_{kij} - s) W_{\tilde{\eta}_{\gamma,k'}}(T_{kil} - t) \right. \\ &\quad \left. \cdot W_{\tilde{\eta}_{\mu,k'}}(T_{k'i'j'} - T_{kil})(Y_{kij} - \mu(T_{kij}))(Y_{k'i'j'} - \mu(T_{k'i'j'})) \right\} \\ &= -S_{K,1}^{-1} S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{j \neq l}^{n_k} \sum_{i'=1}^{m_{ki}} \mathbb{E} \left\{ W_{\tilde{\eta}_{\gamma,k}}(T_{kij} - s) W_{\tilde{\eta}_{\gamma,k}}(T_{kil} - t) W_{\tilde{\eta}_{\mu,k}}(T_{kij'} - T_{kil}) \right. \\ &\quad \left. \cdot (Y_{kij} - \mu(T_{kij}))(Y_{kij'} - \mu(T_{kij'})) \right\} \\ &= -S_{K,1}^{-1} S_{K,2}^{-1} \sum_{k=1}^K \left[\sum_j m_{ki}^2 (m_{ki} - 1) \mathbb{E} \left\{ W_{\tilde{\eta}_{\gamma,k}}(T_{kij} - s) W_{\tilde{\eta}_{\gamma,k}}(T_{kil} - t) \right. \right. \end{aligned}$$

$$\begin{aligned} & \cdot W_{\tilde{\eta}_{\mu,k}}(T_{kij'} - T_{kil})(Y_{kij} - \mu(T_{kij}))(Y_{kij'} - \mu(T_{kij'})) \Big] \\ & = O_p(S_{K,1}^{-1}S_{K,2}^{-1}S_{K,3}), \end{aligned}$$

and hence $E\Delta\tilde{R}_1/Q_{0R_1} = O_p(S_{K,1}^{-1}S_{K,2}^{-1}S_{K,3})$. Note that

$$\begin{aligned} \Delta\tilde{R}_1^2 &= S_{K,1}^{-2}S_{K,2}^{-2} \sum_{j_1 \neq l_1, j_2 \neq l_2} W_{k_1 i_1 j_1}(s)W_{k_1 i_1 l_1}(t)W_{k'_1 i'_1 j'_1}(T_{k_1 i_1 l_1})W_{k_2 i_2 j_2}(s)W_{k_2 i_2 l_2}(t)W_{k'_2 i'_2 j'_2}(T_{k_2 i_2 l_2}) \\ & \cdot (Y_{k_1 i_1 j_1} - \mu_{k_1 i_1 j_1}) \left\{ Y_{k'_1 i'_1 j'_1} - \mu_{k'_1 i'_1 j'_1} - \frac{1}{2}\mu''_{k_1 i_1 l_1}(T_{k'_1 i'_1 j'_1} - T_{k_1 i_1 l_1})^2 \right\} \\ & \cdot (Y_{k_2 i_2 j_2} - \mu_{k_2 i_2 j_2}) \left\{ Y_{k'_2 i'_2 j'_2} - \mu_{k'_2 i'_2 j'_2} - \frac{1}{2}\mu''_{k_2 i_2 l_2}(T_{k'_2 i'_2 j'_2} - T_{k_2 i_2 l_2})^2 \right\}, \end{aligned}$$

$E\Delta\tilde{R}_1^2$ is the summation of the following quantities:

$$\begin{aligned} (a) &= S_{K,1}^{-2}S_{K,2}^{-2}E \sum W_{k_1 i_1 j_1}(s)W_{k_1 i_1 l_1}(t)W_{k'_1 i'_1 j'_1}(T_{k_1 i_1 l_1})W_{k_2 i_2 j_2}(s)W_{k_2 i_2 l_2}(t)W_{k'_2 i'_2 j'_2}(T_{k_2 i_2 l_2}) \\ & \cdot (Y - \mu)_{k_1 i_1 j_1}(Y - \mu)_{k'_1 i'_1 j'_1}(Y - \mu)_{k_2 i_2 j_2}(Y - \mu)_{k'_2 i'_2 j'_2}, \\ (b) &= -\frac{1}{2}S_{K,1}^{-2}S_{K,2}^{-2}E \sum W_{k_1 i_1 j_1}(s)W_{k_1 i_1 l_1}(t)W_{k'_1 i'_1 j'_1}(T_{k_1 i_1 l_1})W_{k_2 i_2 j_2}(s)W_{k_2 i_2 l_2}(t)W_{k'_2 i'_2 j'_2}(T_{k_2 i_2 l_2}) \\ & \cdot (Y - \mu)_{k_1 i_1 j_1}(Y - \mu)_{k'_1 i'_1 j'_1}(Y - \mu)_{k_2 i_2 j_2}\mu''_{k_2 i_2 l_2}(T_{k'_2 i'_2 j'_2} - T_{k_2 i_2 l_2})^2, \\ (c) &= -\frac{1}{2}S_{K,1}^{-2}S_{K,2}^{-2}E \sum W_{k_1 i_1 j_1}(s)W_{k_1 i_1 l_1}(t)W_{k'_1 i'_1 j'_1}(T_{k_1 i_1 l_1})W_{k_2 i_2 j_2}(s)W_{k_2 i_2 l_2}(t)W_{k'_2 i'_2 j'_2}(T_{k_2 i_2 l_2}) \\ & \cdot (Y - \mu)_{k_1 i_1 j_1}(Y - \mu)_{k_2 i_2 j_2}(Y - \mu)_{k'_2 i'_2 j'_2}\mu''_{k_1 i_1 l_1}(T_{k'_1 i'_1 j'_1} - T_{k_1 i_1 l_1})^2, \\ (d) &= \frac{1}{4}S_{K,1}^{-2}S_{K,2}^{-2}E \sum W_{k_1 i_1 j_1}(s)W_{k_1 i_1 l_1}(t)W_{k'_1 i'_1 j'_1}(T_{k_1 i_1 l_1})W_{k_2 i_2 j_2}(s)W_{k_2 i_2 l_2}(t)W_{k'_2 i'_2 j'_2}(T_{k_2 i_2 l_2}) \\ & \cdot (Y - \mu)_{k_1 i_1 j_1}(Y - \mu)_{k_2 i_2 j_2}\mu''_{k_1 i_1 l_1}(T_{k'_1 i'_1 j'_1} - T_{k_1 i_1 l_1})^2\mu''_{k_2 i_2 l_2}(T_{k'_2 i'_2 j'_2} - T_{k_2 i_2 l_2})^2, \end{aligned}$$

which are computed as follows.

- a. For (a), there are four cases according to the relationship between (k_1, i_1) , (k_2, i_2) , (k'_1, i'_1) , (k'_2, i'_2) to make the quantity nonzero:

$$(B.1) \text{ Let } \Lambda_1^a = \{(k_1, i_1) = (k'_1, i'_1), (k_2, i_2) = (k'_2, i'_2), (k_1, i_1) \neq (k_2, i_2), j_1 \neq l_1, j_2 \neq l_2\},$$

$$\begin{aligned} (a.1) &= S_{K,1}^{-2}S_{K,2}^{-2}E \sum_{\Lambda_1^a} W_{k_1 i_1 j_1}(s)W_{k_1 i_1 l_1}(t)W_{k'_1 i'_1 j'_1}(T_{k_1 i_1 l_1})W_{k_2 i_2 j_2}(s)W_{k_2 i_2 l_2}(t) \\ & \cdot W_{k_2 i_2 j'_2}(T_{k_2 i_2 l_2})(Y - \mu)_{k_1 i_1 j_1}(Y - \mu)_{k'_1 i'_1 j'_1}(Y - \mu)_{k_2 i_2 j_2}(Y - \mu)_{k_2 i_2 j'_2} \\ & = S_{K,1}^{-2}S_{K,2}^{-2} \sum_{\Lambda_1^a} E \{ W_{k_1 i_1 l_1}(t)W_{k_1 i_1 j_1}(s)W_{k'_1 i'_1 j'_1}(T_{k_1 i_1 l_1})(Y - \mu)_{k_1 i_1 j_1}(Y - \mu)_{k'_1 i'_1 j'_1} \} \end{aligned}$$

$$\begin{aligned} & \cdot \mathbb{E} \left\{ W_{k_2 i_2 l_2}(t) W_{k_2 i_2 j_2}(s) W_{k_2 i_2 j'_2}(T_{k_2 i_2 l_2})(Y - \mu)_{k_2 i_2 j_2} (Y - \mu)_{k_2 i_2 j'_2} \right\} \\ & = (\mathbb{E} \Delta \tilde{R}_1(s, t))^2. \end{aligned}$$

(B.2) Let $\Lambda_2^a = \{(k_1, i_1) = (k_2, i_2), (k'_1, i'_1) = (k'_2, i'_2), (k_1, i_1) \neq (k'_1, i'_1), j_1 \neq l_1, j_2 \neq l_2, j'_1 \neq j'_2\}$,

$$\begin{aligned} (a.2) & = S_{K,1}^{-2} S_{K,2}^{-2} \sum_{\Lambda_2^a} \mathbb{E} \{ W_{k_1 i_1 j_1}(s) W_{i_2 j_1 k_1}(s) (Y - \mu)_{k_1 i_1 j_1} (Y - \mu)_{i_2 j_1 k_1} \} \\ & \quad \cdot \mathbb{E} \left[W_{k_1 i_1 l_1}(t) W_{l_2 j_1 k_1}(t) W_{k'_1 i'_1 j'_1}(T_{k_1 i_1 l_1}) W_{i'_2 j'_1 k'_1}(T_{l_2 j_1 k_1}) (Y - \mu)_{k'_1 i'_1 j'_1} (Y - \mu)_{i'_2 j'_1 k'_1} \right] \\ & = S_{K,1}^{-2} S_{K,2}^{-2} \sum_{\Lambda_2^a} \gamma(s, s) \gamma(t, t) f^2(s) f^4(t) + O_p(\rho_{\gamma,2} + \rho_{\mu,2}) \\ & = S_{K,1}^{-2} S_{K,2}^{-2} S_{K,3}^2 \{ \gamma(s, s) \gamma(t, t) f^2(s) f^4(t) + O_p(\rho_{\gamma,2} + \rho_{\mu,2}) \}. \end{aligned}$$

(B.3) Let $\Lambda_3^a = \{(k_1, i_1) = (k'_2, i'_2), (k_2, i_2) = (k'_1, i'_1), (k_1, i_1) \neq (k_2, i_2), j_1 \neq l_1, j_2 \neq l_2\}$,

$$\begin{aligned} (a.3) & = S_{K,1}^{-2} S_{K,2}^{-2} \mathbb{E} \left[\sum_{\Lambda_3^a} W_{k_1 i_1 j_1}(s) W_{k_1 i_1 l_1}(t) W_{k_2 i_2 j'_1}(T_{k_1 i_1 l_1}) \right. \\ & \quad \cdot W_{k_2 i_2 j_2}(s) W_{k_2 i_2 l_2}(t) W_{k_1 i_1 j'_2}(T_{k_2 i_2 l_2}) \\ & \quad \left. \cdot (Y - \mu)_{k_1 i_1 j_1} (Y - \mu)_{k_2 i_2 j'_1} (Y - \mu)_{k_2 i_2 j_2} (Y - \mu)_{k_1 i_1 j'_2} \right] \\ & = (\mathbb{E} \Delta \tilde{R}_1(s, t))^2. \end{aligned}$$

(B.4) Let $\Lambda_4^a = \{(k_1, i_1) = (k_2, i_2) = (k'_1, i'_1) = (k'_2, i'_2), j_1 \neq l_1, j_2 \neq l_2\}$ and omit the subscript ki ,

$$\begin{aligned} (a.4) & = S_{K,1}^{-2} S_{K,2}^{-2} \mathbb{E} \left\{ \sum_{\Lambda_4^a} W_{j_1}(s) W_{l_1}(t) W_{j'_1}(T_{l_1}) W_{j_2}(s) W_{l_2}(t) W_{j'_2}(T_{l_2}) \right. \\ & \quad \left. \cdot (Y - \mu)_{j_1} (Y - \mu)_{j'_1} (Y - \mu)_{j_2} (Y - \mu)_{j'_2} \right\}. \end{aligned}$$

It is the summation of the following sets,

$$\Lambda_{4.1}^a = \{j_1 = j_2, j'_1 = j'_2, l_1 = l_2, j_1 \neq l_1, j_2 \neq l_2\},$$

$$\Lambda_{4.2}^a = \{j_1 = j_2, j'_1 \neq j'_2, j_1 \neq l_1, j_2 \neq l_2\} \cup \{j_1 \neq j_2, j'_1 = j'_2, l_1 = l_2, j_1 \neq l_1, j_2 \neq l_2\},$$

$$\Lambda_{4.3}^a = \{j_1, j_2, j'_1, j'_2 \text{ are not equal}, j_1 \neq l_1, j_2 \neq l_2\}.$$

Correspondingly, we define

$$\begin{aligned}
E_{1|4}(s, t) &= \mathbb{E}\{(Y_1 - \mu(T_1))^2(Y_2 - \mu(T_2))^2|T_1 = s, T_2 = t\}, \\
E_{2|4}(s, t) &= \mathbb{E}\{(Y_1 - \mu(T_1))^2(Y_2 - \mu(T_2))(Y_3 - \mu(T_3))|T_1 = s, T_2 = t, T_3 = t\}, \\
E_{3|4}(s, t) &= \mathbb{E}\{(Y_1 - \mu(T_1))(Y_2 - \mu(T_2))(Y_3 - \mu(T_3))(Y_4 - \mu(T_4))|T_1 = s, \\
&\quad T_2 = t, T_3 = s, T_4 = t\}.
\end{aligned}$$

Then by easy calculation,

$$\begin{aligned}
(a.4)|_{\Lambda_{4.1}^a} &\doteq S_{K,1}^{-2}S_{K,2}^{-2}\left(\sum_{k=1}^K s_{k,3}\tilde{\eta}_{\mu,k}^{-1}\tilde{\eta}_{\gamma,k}^{-2}\right)R(W)^3E_{1|4}(s, t)f(s)f^2(t), \\
(a.4)|_{\Lambda_{4.2}^a} &\doteq S_{K,1}^{-2}S_{K,2}^{-2}\left\{\left(\sum_{k=1}^K s_{k,5}\tilde{\eta}_{\gamma,k}^{-1}\right)R(W)E_{2|4}(s, t)f(s)f^4(t)\right. \\
&\quad \left.+\left(\sum_{k=1}^K s_{k,4}\tilde{\eta}_{\mu,k}^{-1}\tilde{\eta}_{\gamma,k}^{-1}\right)R(W)^2E_{2|4}(t, s)f^2(s)f^2(t)\right\}, \\
(a.4)|_{\Lambda_{4.3}^a} &\doteq S_{K,1}^{-2}S_{K,2}^{-2}S_6E_{3|4}(s, t)f^2(s)f^4(t).
\end{aligned}$$

b. For (b) and (c), it is required that $(k_1, i_1) = (k_2, i_2) = (k'_1, i'_1)$, define

$$E_{2|3}(s, s, t) = \mathbb{E}\{(Y_1 - \mu(T_1))(Y_2 - \mu(T_2))(Y_3 - \mu(T_3))|T_1 = s, T_2 = s, T_3 = t\}.$$

By easy calculation,

$$(b) + (c) \doteq S_{K,1}^{-2}S_{K,2}^{-2}\alpha(W)E_{2|3}(s, s, t)\left(\sum_{k=1}^K s_{k,3}\tilde{\eta}_{\mu,k}^2\right).$$

c. For (d), we need $(k_1, i_1) = (k_2, i_2)$, then

$$(d) \doteq \frac{1}{4}S_{K,2}^{-2}\alpha^2(W)\{\mu''(t)\}^2\gamma(s, s)f^2(s)f^4(t)\left(\sum_{k=1}^K s_{k,4}\tilde{\eta}_{\gamma,k}^4\right).$$

From the above arguments,

$$\begin{aligned}
\text{Var } \Delta\tilde{R}_1 &= \mathbb{E}\Delta\tilde{R}_1^2 - (\mathbb{E}\Delta\tilde{R}_1)^2 \\
&= O_p\left(S_{K,1}^{-2}S_{K,2}^{-2}\left\{S_{K,3}^2 + \sum_k s_{k,3}\tilde{\eta}_{\mu,k}^{-1}\tilde{\eta}_{\gamma,k}^{-2} + \sum_k s_{k,4}\tilde{\eta}_{\mu,k}^{-1}\tilde{\eta}_{\gamma,k}^{-1} + \sum_k s_{k,5}\tilde{\eta}_{\gamma,k}^{-1}\right\}\right. \\
&\quad \left. + S_{K,2}^{-2}\sum_k s_{k,4}\tilde{\eta}_{\gamma,k}^4\right).
\end{aligned}$$

Note that

$$\begin{aligned}\text{Cov}(Q_{0R_1}, \Delta \tilde{R}_1) &= -S_{K,1}^{-2} S_{K,2}^{-1} \left(\sum_{k=1}^K \sum_{i=1}^{n_k} m_{ki}^2 (m_{ki} - 1) \right) \rho_{\mu,-1} R(W) \gamma(s, t) f(s) f^3(t), \\ \text{E}(Q_{0R_1}) &= f(t) + \frac{1}{2} \alpha(W) f''(t) \rho_{\mu,2} + o_p(\rho_{\mu,2}), \\ \text{Var}(Q_{0R_1}) &= \left(\frac{\rho_{\mu,-1}}{S_{K,1}} \right) R(W) f(t) + o_p \left(\frac{\rho_{\mu,-1}}{S_{K,1}} \right).\end{aligned}$$

By the delta method,

$$\begin{aligned}\text{Var} \left(\frac{\Delta \tilde{R}_1}{Q_{0R_1}} \right) &= \frac{\text{Var} \Delta \tilde{R}_1}{\text{E}^2 Q_{0R_1}} - 2 \frac{\text{E} \Delta \tilde{R}_1 \text{Cov}(Q_{0R_1}, \Delta \tilde{R}_1)}{\text{E}^3 Q_{0R_1}} + \frac{\text{E}^2 \Delta \tilde{R}_1 \text{Var} Q_{0R_1}}{\text{E}^4 Q_{0R_1}} \\ &= O_p \left(S_{K,1}^{-2} S_{K,2}^{-2} \left\{ S_{K,3}^2 + \sum_k s_{k,3} \tilde{\eta}_{\mu,k}^{-1} \tilde{\eta}_{\gamma,k}^{-2} + \sum_k s_{k,4} \tilde{\eta}_{\mu,k}^{-1} \tilde{\eta}_{\gamma,k}^{-1} + \sum_k s_{k,5} \tilde{\eta}_{\gamma,k}^{-1} \right\} \right. \\ &\quad \left. + S_{K,2}^{-2} \sum_k s_{k,4} \tilde{\eta}_{\gamma,k}^4 \right).\end{aligned}$$

■

Now we give the proof of the asymptotic distribution of $\tilde{\gamma}^{(K)}$ in Theorem 2.

Proof. Denote

$$f_1 = Q_{20}Q_{02} - Q_{11}^2, \quad f_2 = Q_{10}Q_{02} - Q_{01}Q_{11}, \quad f_3 = Q_{01}Q_{20} - Q_{10}Q_{11}.$$

Let $\tilde{\gamma}$ be the local linear estimate based on $\tilde{C}_{kijl} = (Y_{kij} - \tilde{\mu}_{kij})(Y_{kil} - \tilde{\mu}_{kil})$ which can be expressed as

$$\tilde{\gamma} = \frac{f_1 R_{00} - f_2 R_{10} - f_3 R_{01}}{f_1 Q_{00} - f_2 Q_{10} - f_3 Q_{01}},$$

where for $p, q = 0, 1, 2$,

$$\begin{aligned}Q_{p,q} &= S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{1 \leq j \neq l \leq m_{ki}} W_{\tilde{\eta}_{\gamma,k}}(T_{kij} - s) W_{\tilde{\eta}_{\gamma,k}}(T_{kil} - t) \left(\frac{T_{kij} - s}{\tilde{\eta}_{\gamma,k}} \right)^p \left(\frac{T_{kil} - t}{\tilde{\eta}_{\gamma,k}} \right)^q \\ &= \sum_k \omega_k Q_{pq,k}, \\ R_{p,q} &= S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{1 \leq j \neq l \leq m_{ki}} W_{\tilde{\eta}_{\gamma,k}}(T_{kij} - s) W_{\tilde{\eta}_{\gamma,k}}(T_{kil} - t) \left(\frac{T_{kij} - s}{\tilde{\eta}_{\gamma,k}} \right)^p \left(\frac{T_{kil} - t}{\tilde{\eta}_{\gamma,k}} \right)^q \tilde{C}_{kijl} \\ &= \sum_k \omega_k R_{pq,k}.\end{aligned}$$

Let

$$\begin{aligned}
\tilde{\beta}_{1,k} &= \frac{1}{\tilde{\eta}_{\gamma,k}} \frac{f_2(R_{10} - R_{00}Q_{10}/Q_{00})}{f_1Q_{10} - f_2Q_{10}^2/Q_{00} - f_3Q_{01}Q_{10}/Q_{00}}, \\
\tilde{\beta}_{2,k} &= \frac{1}{\tilde{\eta}_{\gamma,k}} \frac{f_3(R_{01} - R_{00}Q_{01}/Q_{00})}{f_1Q_{01} - f_2Q_{01}Q_{10}/Q_{00} - f_3Q_{01}^2/Q_{00}}, \\
\tilde{\beta}_{1,k}^* &= \frac{1}{\tilde{\eta}_{\gamma,k}} \frac{f_2(R_{10}^* - R_{00}^*Q_{10}/Q_{00})}{f_1Q_{10} - f_2Q_{10}^2/Q_{00} - f_3Q_{01}Q_{10}/Q_{00}}, \\
\tilde{\beta}_{2,k}^* &= \frac{1}{\tilde{\eta}_{\gamma,k}} \frac{f_3(R_{01}^* - R_{00}^*Q_{01}/Q_{00})}{f_1Q_{01} - f_2Q_{01}Q_{10}/Q_{00} - f_3Q_{01}^2/Q_{00}}, \\
R_{p,q}^* &= S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{1 \leq j \neq l \leq m_{ki}} W_{\tilde{\eta}_{\gamma,k}}(T_{kij} - s) W_{\tilde{\eta}_{\gamma,k}}(T_{kil} - t) \left(\frac{T_{kij} - s}{\tilde{\eta}_{\gamma,k}} \right)^p \left(\frac{T_{kil} - t}{\tilde{\eta}_{\gamma,k}} \right)^q C_{kijl}, \\
\tilde{G}_{pq} &= S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{1 \leq j \neq l \leq m_{ki}} W_{\tilde{\eta}_{\gamma,k}}(T_{kij} - s) W_{\tilde{\eta}_{\gamma,k}}(T_{kil} - t) T_{kij}^p T_{kil}^q = \sum_k \omega_k \tilde{G}_{pq,k}.
\end{aligned}$$

Then we can write

$$\begin{aligned}
\tilde{\gamma} &= \frac{R_{00}}{Q_{00}} - \sum_k \omega_k \tilde{\beta}_{1,k} \cdot \frac{\tilde{G}_{10,k} - sQ_{00}}{Q_{00}} - \sum_k \omega_k \tilde{\beta}_{2,k} \cdot \frac{\tilde{G}_{01,k} - tQ_{00}}{Q_{00}}, \\
\bar{\gamma} &= \frac{R_{00}}{Q_{00}} - \frac{\partial \gamma}{\partial s} \cdot \frac{\tilde{G}_{10} - sQ_{00}}{Q_{00}} - \frac{\partial \gamma}{\partial t} \cdot \frac{\tilde{G}_{01} - tQ_{00}}{Q_{00}}, \\
\bar{\gamma}^* &= \frac{R_{00}^*}{Q_{00}} - \frac{\partial \gamma}{\partial s} \cdot \frac{\tilde{G}_{10} - sQ_{00}}{Q_{00}} - \frac{\partial \gamma}{\partial t} \cdot \frac{\tilde{G}_{01} - tQ_{00}}{Q_{00}}.
\end{aligned}$$

Following similar arguments in the proof of Theorem 1, we can prove that

$$\tilde{\gamma} - \bar{\gamma} = o_p \left[\min \left\{ \left(S_{K,2}^{-2} \sum_{k=1}^K \sum_{i=1}^{n_k} \frac{S_{k,3}}{\tilde{\eta}_{\gamma,k}} \right)^{\frac{1}{2}}, (S_{K,4}/S_{K,2}^2)^{\frac{1}{2}} \right\} \right].$$

Now we derive the asymptotic distribution of $\bar{\gamma}$. We first show that $\bar{\gamma} - \bar{\gamma}^* = O_p\{(N_K)^{-1}\}$.

(B.1) Calculate $\bar{\gamma} - \bar{\gamma}^*$.

From the definition of $\bar{\gamma}$ and $\bar{\gamma}^*$, we have $\bar{\gamma} - \bar{\gamma}^* = R_{00}/Q_{00}$. Note that

$$\begin{aligned}
R_{00} &= S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{j \neq l} W_{kij}(s) W_{kil}(t) \tilde{C}_{kijl} \\
&= S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{j \neq l} W_{kij}(s) W_{kil}(t) \{ (Y_{kij} - \mu_{kij})(Y_{kil} - \mu_{kil}) + (Y_{kij} - \mu_{kij})(\mu_{kil} - \tilde{\mu}_{kil}) \\
&\quad + (\mu_{kij} - \tilde{\mu}_{kil})(Y_{kil} - \mu_{kil}) + (\tilde{\mu}_{kij} - \mu_{kij})(\tilde{\mu}_{kil} - \mu_{kil}) \}
\end{aligned}$$

$$\begin{aligned} & \doteq R_{00}^* + S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{j \neq l} W_{kij}(s) W_{kil}(t) \{ (Y_{kij} - \mu_{kij})(\mu_{kil} - \tilde{\mu}_{kil}) \\ & \qquad \qquad \qquad + (\mu_{kij} - \tilde{\mu}_{kil})(Y_{kil} - \mu_{kil}) \}. \end{aligned}$$

Recall that in S.1, we have

$$\tilde{\mu}(t) - \bar{\mu}(t) = o_p(\bar{\mu}(t) - \mu(t)),$$

where

$$\begin{aligned} \bar{\mu}(t) &= \frac{R_0}{Q_0} - \frac{\check{G}_1}{Q_0} \mu'(t) + t \frac{\check{G}_0}{Q_0} \mu'(t), \\ Q_r(t) &= \frac{1}{S_{K,1}} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{j=1}^{m_{ki}} W_{\tilde{\eta}_{\mu,k}}(T_{kij} - t) \left(\frac{T_{kij} - t}{\tilde{\eta}_{\mu,k}} \right)^r, \\ R_r(t) &= \frac{1}{S_{K,1}} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{j=1}^{m_{ki}} W_{\tilde{\eta}_{\mu,k}}(T_{kij} - t) \left(\frac{T_{kij} - t}{\tilde{\eta}_{\mu,k}} \right)^r Y_{kij}, \\ \check{G}_r(t) &= \frac{1}{S_{K,1}} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{j=1}^{m_{ki}} W_{\tilde{\eta}_{\mu,k}}(T_{kij} - t) T_{kij}^r. \end{aligned}$$

then one can obtain

$$\begin{aligned} & S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{j \neq l} W_{kij}(s) W_{kil}(t) (Y_{kij} - \mu_{kij})(\mu_{kil} - \tilde{\mu}_{kil}) \\ & \doteq S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{j \neq l} W_{kij}(s) W_{kil}(t) (Y_{kij} - \mu_{kij})(\mu_{kil} - \bar{\mu}_{kil}) \\ & = S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{j \neq l} W_{kij}(s) W_{kil}(t) (Y_{kij} - \mu_{kij}) \left\{ \mu(T_{kil}) - \frac{R_0}{Q_0} + \frac{\check{G}_1}{Q_0} \mu'(T_{kil}) - T_{kil} \frac{\check{G}_0}{Q_0} \mu'(T_{kil}) \right\} \\ & = S_{K,2}^{-1} Q_0^{-1} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{j \neq l} W_{kij}(s) W_{kil}(t) (Y_{kij} - \mu_{kij}) \\ & \quad \cdot \left\{ Q_0(T_{kil}) \mu(T_{kil}) - R_0(T_{kil}) + \check{G}_1(T_{kil}) \mu'(T_{kil}) - T_{kil} \check{G}_0(T_{kil}) \mu'(T_{kil}) \right\} \\ & = S_{K,1}^{-1} S_{K,2}^{-1} Q_0^{-1} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{j \neq l} \sum_{k'=1}^K \sum_{i'=1}^{n_{k'}} \sum_{j'=1}^{m_{k'i'}} W_{\tilde{\eta}_{\mu,k'}}(T_{kij} - s) W_{\tilde{\eta}_{\mu,k'}}(T_{kil} - t) W_{\tilde{\eta}_{\mu,k'}}(T_{k'i'j'} - T_{kil}) \\ & \quad \cdot (Y_{kij} - \mu(T_{kij})) \{ \mu(T_{kil}) - Y_{k'i'j'} + T_{k'i'j'} \mu'(T_{kil}) - T_{kil} \mu'(T_{kil}) \}. \end{aligned}$$

Take the Taylor expansion of $\mu(T_{kil})$ at $\mu(T_{k'i'j'})$ in the last bracket,

$$S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{j \neq l} W_{kij}(s) W_{kil}(t) (Y_{kij} - \mu_{kij})(\mu_{kil} - \tilde{\mu}_{kil})$$

$$\begin{aligned}
& \doteq - S_{K,1}^{-1} S_{K,2}^{-1} Q_0^{-1} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{j \neq l} \sum_{k'=1}^K \sum_{i'=1}^{n_{k'}} \sum_{j'=1}^{m_{k'i'}} W_{\tilde{\eta}_{\gamma,k'}}(T_{kij} - s) W_{\tilde{\eta}_{\gamma,k'}}(T_{kil} - t) \\
& \quad \cdot W_{\tilde{\eta}_{\mu,k'}}(T_{k'i'j'} - T_{kil})(Y_{kij} - \mu(T_{kij})) \left\{ Y_{k'i'j'} - \mu(T_{k'i'j'}) + \right. \\
& \quad \left. \cdot (\mu'(T_{kil}) - \mu'(T_{k'i'j'}))(T_{k'i'j'} - T_{kil}) + \frac{1}{2} \mu''(T_{kil})(T_{k'i'j'} - T_{kil})^2 \right\}.
\end{aligned}$$

Then take the Taylor expansion of $\mu'(T_{k'i'j'})$ at $\mu'(T_{kil})$ in the last bracket,

$$\begin{aligned}
& S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{j \neq l} W_{kij}(s) W_{kil}(t) (Y_{kij} - \mu_{kij})(\mu_{kil} - \tilde{\mu}_{kil}) \\
& \doteq - S_{K,1}^{-1} S_{K,2}^{-1} Q_0^{-1} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{j \neq l} \sum_{k'=1}^K \sum_{i'=1}^{n_{k'}} \sum_{j'=1}^{m_{k'i'}} W_{\tilde{\eta}_{\gamma,k'}}(T_{kij} - s) W_{\tilde{\eta}_{\gamma,k'}}(T_{kil} - t) \\
& \quad \cdot W_{\tilde{\eta}_{\mu,k'}}(T_{k'i'j'} - T_{kil})(Y_{kij} - \mu(T_{kij})) \left\{ Y_{k'i'j'} - \mu(T_{k'i'j'}) - \frac{1}{2} \mu''(T_{kil})(T_{k'i'j'} - T_{kil})^2 \right\}.
\end{aligned}$$

Denote

$$\begin{aligned}
\Delta \tilde{R}_1 &= - S_{K,1}^{-1} S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{j \neq l} \sum_{k'=1}^K \sum_{i'=1}^{n_{k'}} \sum_{j'=1}^{m_{k'i'}} W_{\tilde{\eta}_{\gamma,k'}}(T_{kij} - s) W_{\tilde{\eta}_{\gamma,k'}}(T_{kil} - t) \\
& \quad \cdot W_{\tilde{\eta}_{\mu,k'}}(T_{k'i'j'} - T_{kil})(Y_{kij} - \mu(T_{kij})) \left\{ Y_{k'i'j'} - \mu(T_{k'i'j'}) - \frac{1}{2} \mu''(T_{kil})(T_{k'i'j'} - T_{kil})^2 \right\}, \\
\Delta \tilde{R}_2 &= - S_{K,1}^{-1} S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^{n_k} \sum_{j \neq l} \sum_{k'=1}^K \sum_{i'=1}^{n_{k'}} \sum_{j'=1}^{m_{k'i'}} W_{\tilde{\eta}_{\gamma,k'}}(T_{kij} - s) W_{\tilde{\eta}_{\gamma,k'}}(T_{kil} - t) \\
& \quad \cdot W_{\tilde{\eta}_{\mu,k'}}(T_{k'i'j'} - T_{kij})(Y_{kil} - \mu(T_{kil})) \left\{ Y_{k'i'j'} - \mu(T_{k'i'j'}) - \frac{1}{2} \mu''(T_{kij})(T_{k'i'j'} - T_{kij})^2 \right\},
\end{aligned}$$

$$Q_{0R1} = Q_0(T_{kil}), \quad Q_{0R2} = Q_0(T_{kij}).$$

It can be written as

$$R_{00} \doteq R_{00}^* + \frac{\Delta \tilde{R}_1}{Q_{0R1}} + \frac{\Delta \tilde{R}_2}{Q_{0R2}},$$

The last terms in the right hand side are of the same order. By Lemma S.1, we have

$$\begin{aligned}
\mathbb{E}(\Delta \tilde{R}_1 / Q_{0R1}) &= O_p(S_{K,1}^{-1} S_{K,2}^{-1} S_{K,3}), \\
\text{Var}(\Delta \tilde{R}_1 / Q_{0R1}) &= O_p \left(S_{K,1}^{-2} S_{K,2}^{-2} \left\{ S_{K,3}^2 (1 + S_{K,1}^{-1} \rho_{\mu,-1}) + S_{K,3} \rho_{\mu,-1} \rho_{\gamma,-2} \right. \right. \\
& \quad \left. \left. + S_4 \rho_{\mu,-1} \rho_{\gamma,-1} + S_5 \rho_{\gamma,-1} + S_6 \right\} + S_{K,2}^{-2} S_4 (\rho_{\mu,2})^2 \right).
\end{aligned}$$

Hence

$$\begin{aligned}
\bar{\gamma} - \bar{\gamma}^* &= O_p \left(S_{K,1}^{-1} S_{K,2}^{-1} \left\{ S_{K,3}^2 (1 + S_{K,1}^{-1} \rho_{\mu,-1}) + S_{K,3} \rho_{\mu,-1} \rho_{\gamma,-2} \right. \right. \\
& \quad \left. \left. + S_{K,4} \rho_{\mu,-1} \rho_{\mu,-1} + S_5 \rho_{\gamma,-1} + S_6 \right\}^{\frac{1}{2}} + S_{K,2}^{-1} S_4^{\frac{1}{2}} (\rho_{\mu,2}) \right).
\end{aligned}$$

(B.2) *Asymptotic distribution of $\bar{\gamma}^*$.*

Define

$$\bar{\gamma}^* = \frac{R_{00}^*}{Q_{00}} - \frac{\partial \gamma}{\partial s} \cdot \frac{\tilde{G}_{10} - sQ_{00}}{Q_{00}} - \frac{\partial \gamma}{\partial t} \cdot \frac{\tilde{G}_{01} - tQ_{00}}{Q_{00}}.$$

The asymptotic normality of $\bar{\gamma}^*$ is accomplished by the asymptotic joint normality of $(R_{00}^* - ER_{00}^*, \tilde{G}_{00} - E\tilde{G}_{00}, \tilde{G}_{10} - E\tilde{G}_{10}, \tilde{G}_{01} - E\tilde{G}_{01})$. Specifically, let

$$E_1(s, t) = E\{(Y_1 - \mu(T_1))^2(Y_2 - \mu(T_2))^2 | T_1 = s, T_2 = t\},$$

$$E_2(s, t) = E\{(Y_1 - \mu(T_1))^2(Y_2 - \mu(T_2))(Y_2 - \mu(T_3)) | T_1 = s, T_2 = t, T_3 = t\}.$$

By easy calculation, we have

$$\begin{aligned} E\tilde{G}_{pq} &= s^p t^q f(s) f(t) + \frac{1}{2} \alpha(W) \{s^p t^q f''(s) f(t) + s^p t^q f(s) f''(t)\} \cdot \bar{\eta}_\gamma^2 \\ &\quad + \frac{1}{2} \alpha(W) \{2p f'(s) f(t) + 2q f'(t) f(s)\} \cdot \bar{\eta}_\gamma^2 + o_p\left(\bar{\eta}_\gamma^2\right), \\ ER_{00}^* &= \gamma(s, t) f(s) f(t) + \frac{1}{2} \alpha(W) \bar{\eta}_\gamma^2 \left\{ \frac{\partial^2 \gamma}{\partial s^2}(s, t) f(s) f(t) \right. \\ &\quad + 2 \frac{\partial \gamma}{\partial s}(s, t) f'(s) f(t) + \gamma(s, t) f''(s) f(t) + \frac{\partial^2 \gamma}{\partial t^2}(s, t) f(t) f(s) \\ &\quad \left. + 2 \frac{\partial \gamma}{\partial t}(s, t) f'(t) f(s) + \gamma(s, t) f''(t) f(s) \right\} + o_p\left(\bar{\eta}_\gamma^2\right), \end{aligned}$$

$$\begin{aligned} \text{Var}(R_{00}^*) &= \{1 + I(s = t)\} \left[R(W)^2 E_1(s, t) f(s) f(t) S_{K,2}^{-2} \bar{\eta}_\gamma^{-2} \right. \\ &\quad \left. + R(W) \left\{ f^2(s) f(t) E_2(t, s) + f(s) f(t)^2 E_2(s, t) \right\} S_{K,2}^{-1} \sum_{k=1}^K \frac{s_{k,3}}{\tilde{\eta}_{\gamma,k}} \right] \\ &\quad + V_3(s, t) f^2(s) f^2(t) S_{K,2}^{-2} S_4 + o\left(S_{K,2}^{-2} \bar{\eta}_\gamma^{-2}\right) \\ &\quad + o_p\left(S_{K,2}^{-2} \sum_{k=1}^K \frac{s_{k,3}}{\tilde{\eta}_{\gamma,k}}\right) + o_p\left(S_{K,2}^{-2} S_4\right), \end{aligned}$$

$$\begin{aligned} \text{Cov}(\tilde{G}_{pq}, \tilde{G}_{p'q'}) &= \{1 + I(s = t)\} s^{p+p'} t^{q+q'} \left[R(W)^2 f(s) f(t) S_{K,2}^{-2} \bar{\eta}_\gamma^{-2} \right. \\ &\quad \left. + R(W) \{f^2(s) f(t) + f(s) f^2(t)\} S_{K,2}^{-2} \sum_{k=1}^K \frac{s_{k,3}}{\tilde{\eta}_{\gamma,k}} \right] \\ &\quad + o_p\left(S_{K,2}^{-2} \bar{\eta}_\gamma^{-2}\right) + o_p\left(S_{K,2}^{-2} \sum_{k=1}^K \frac{s_{k,3}}{\tilde{\eta}_{\gamma,k}}\right), \end{aligned}$$

$$\text{Cov}(\tilde{G}_{pq}, R_{00}^*) = \{1 + I(s = t)\} s^p t^q \left[R(W)^2 \gamma(s, t) f(s) f(t) S_{K,2}^{-2} \bar{\eta}_\gamma^{-2} \right.$$

$$\begin{aligned}
& + R(W)\gamma(s, t)\{f^2(s)f(t) + f(s)f^2(t)\}S_{K,2}^{-2} \left[\sum_{k=1}^K \frac{s_{k,3}}{\tilde{\eta}_{\gamma,k}} \right] \\
& + o_p \left(S_{K,2}^{-2} \overline{\eta_\gamma^2} \right) + o_p \left(S_{K,2}^{-2} \sum_{k=1}^K \frac{s_{k,3}}{\tilde{\eta}_{\gamma,k}} \right).
\end{aligned}$$

From the delta method, $\bar{\gamma}(t)$ follows the asymptotic normality that

$$\Gamma_{\bar{\gamma}^*}^{-1/2} \left\{ \bar{\gamma}^*(s, t) - \gamma(s, t) - \frac{1}{2} \overline{\eta_\gamma^2} \alpha(W) \left(\frac{\partial^2 \gamma}{\partial s^2}(s, t) + \frac{\partial^2 \gamma}{\partial t^2}(s, t) \right) + o_p \left(\overline{\eta_\gamma^2} \right) \right\} \xrightarrow{d} N(0, 1),$$

where

$$\begin{aligned}
\Gamma_{\bar{\gamma}^*} = & \{1 + I(s = t)\} \left\{ S_{K,2}^{-1} \overline{\eta_\gamma^2} R(W)^2 \frac{V_1(s, t)}{f(s)f(t)} \right. \\
& \left. + S_{K,2}^{-2} \sum_{k=1}^K \frac{s_{k,3}}{\tilde{\eta}_{\gamma,k}} R(W) \frac{f(s)V_2(t, s) + f(t)V_2(s, t)}{f(s)f(t)} \right\} + V_3(s, t) S_4 / S_{K,2}^2.
\end{aligned}$$

The proof is completed combining the above arguments. ■

S.3. PROOF OF THEOREM 3

We mention that Theorem 3 and Theorem 4 can be extended to the standard d -dimensional local linear regression. To see this, we introduce the following notations and assumptions.

Notations and assumptions for d -dimensional local linear regression. The underlying model is $Y = m(X) + \varepsilon$ where $X \in \mathcal{R}^d$, $E\varepsilon = 0$ and $\text{Var } \varepsilon = \sigma^2(X)$. Suppose we observe $\{(Y_{ki}, X_{ki} : i = 1, \dots, n_k)\}$ in the k th block and $N_K = \sum_{k=1}^K n_k$, i.e., we use n_k to denote the sub-sample size of the K th data block and $N_K = \sum_{k=1}^K n_k$ to denote the full sample size up to time K , which correspond to $s_{K,1}, S_{K,1}$ in (6) for mean estimation and $s_{K,2}, S_{K,2}$ for covariance estimation, respectively. We impose the following assumptions which are in parallel to (A.1)–(A.3) and (A.6).

(B.1) Observations $\{X_{ki} : i = 1, \dots, n_k, k = 1, \dots, K\}$ are i.i.d. copies of a random variable X defined on $[0, 1]^d$ whose density $f(\cdot)$ is bounded away from 0 with bounded second derivative.

(B.2) The second and fourth derivatives of regression function $m(x)$ are bounded and continuous.

(B.3) Noises ε_{Ki} are i.i.d. with mean 0 and variance $\sigma^2(x)$.

(B.4) The block size $Kn_k/N_K \rightarrow 1$ as $K \rightarrow \infty$ for $k = 1, 2, \dots, K$.

The optimal bandwidth at time K is $h_*^{(K)} \asymp N_K^{-1/(4+d)}$. The candidate sequence is $\{\eta_l^{(K)}\}_{l=1}^L$ and the pseudo-bandwidths are $\{\tilde{\eta}_k^{(K)} : k = 1, \dots, K\}$ for $j = 1, \dots, d$. The centroids for the candidate sequence are $\{\phi_i^{(K)} : i = 1, \dots, L\}$. Further define

$$\rho_i^{(K)} = \sum_{k=1}^K \frac{n_k}{N_K} \left(\tilde{\eta}_k^{(K)} \right)^i.$$

We also mention that the mean and covariance estimates corresponds the cases $d = 1$ and $d = 2$, respectively.

When estimating μ , all observations are pooled together and the index of subjects become invalid, which is intrinsically one-dimensional local linear regression. Hence one can derive the following result for standard d -dimensional local linear regression by the same arguments in Theorem 1.

Lemma S. 2. *Suppose assumptions (B.1)–(B.3) and (A.8)–(A.9) hold. If*

$$\tilde{h}^{(K)} - h_*^{(K)} = o_p \left(N_K^{-\frac{1}{4+d}} \right), \quad j = 1, \dots, d, \quad (\text{S.3})$$

then a fixed interior point $x \in (0, 1)^d$, as $K \rightarrow \infty$, the online estimate $\tilde{m}^{(K)}(x)$ satisfies

$$\left\{ N_K / \rho_{-d}^{(K)} \right\}^{\frac{1}{2}} \left[\tilde{m}^{(K)}(x) - m(x) - \frac{1}{2} \alpha(W) \left\{ \sum_{j=1}^d m_j''(x) \right\} \rho_2 + o_p(\rho_2) \right] \xrightarrow{d} N \left(0, \frac{R(W)^d \sigma^2(x)}{f(x)} \right).$$

We prove in Section S.3 that (S.3) is satisfied by the proposed online method.

Denote $\mathbb{X} = \{X_{ki} \in \mathcal{R}^d : k = 1, \dots, K; i = 1, \dots, n_k\}$, $\theta = \int \{\sum_{j=1}^d \partial^2 m(x) / \partial x_j^2\} f(x) dx$ and $\nu = \int_{[0,1]^d} R(W)^d \sigma^2(x) dx$. Similarly, we adopt another online local cubic and online local linear to estimate θ and ν , which are denoted as $\tilde{\theta}^{(K)}$ and $\tilde{\nu}^{(K)}$, respectively. Then the optimal bandwidth and the online estimated bandwidth for $m(\cdot)$ are

$$h_*^{(K)} = \left(\frac{\nu}{\alpha^2(W)\theta} \right)^{\frac{1}{4+d}} N_K^{-\frac{1}{4+d}}, \quad \tilde{h}^{(K)} = \left(\frac{\tilde{\nu}^{(K)}}{\alpha^2(W)\tilde{\theta}^{(K)}} \right)^{\frac{1}{4+d}} N_K^{-\frac{1}{4+d}}. \quad (\text{S.4})$$

The candidate sequence for θ and ν are $\{\eta_{\theta,l}^{(K)}\}_{l=1}^J$ and $\{\eta_{\nu,l}^{(K)}\}_{l=1}^J$, respectively, and the pseudo-bandwidths are $\{\tilde{\eta}_{\theta,k}^{(K)} : k = 1, \dots, K\}$ and $\{\tilde{\eta}_{\nu,k}^{(K)} : k = 1, \dots, K\}$. Then we have the following conclusion.

Lemma S. 3. *Suppose assumptions (B.1)–(B.4) and (A.8)–(A.9) hold. When the online bandwidths and candidate sequence for θ and ν are*

$$h_\theta^{(K)} = GN_K^{-1/(6+d)}, \quad h_\nu^{(K)} = RN_K^{-1/(4+d)}, \quad 1 < G, R < \infty,$$

$$\eta_{\theta,J}^{(K)} < \dots < \eta_{\theta,1}^{(K)} = h_\theta^{(K)}, \quad \eta_{\nu,J}^{(K)} < \dots < \eta_{\nu,1}^{(K)} = h_\nu^{(K)},$$

then as $K \rightarrow \infty$, $\tilde{h}^{(K)}$ in (S.4) satisfies

$$\frac{\tilde{h}^{(K)} - h_*^{(K)}}{h_*^{(K)}} = O_p\left(N_K^{-\frac{2}{d+2}}\right).$$

Proof. Using the same technique in the proof S.1, we can derive that

$$\begin{aligned} \mathbb{E}\{(\tilde{\theta}^{(K)} - \theta)^2 \mid \mathbb{X}\} &= (A_1\rho_{\theta,2} + A_2N_K^{-1}\rho_{\theta,-4-d})^2 + A_3N_K^{-2}\rho_{\theta,-8-d} \\ &\quad + o_p\left(\{\rho_{\theta,2} + N_K^{-1}\rho_{\theta,-4-d}\}^2 + N_K^{-2}\rho_{\theta,-8-d}\right), \end{aligned}$$

where A_1, A_2, A_3 are the same constants as in Wand and Jones (1994), and

$$\rho_{\theta,i} = \frac{1}{N_K} \sum_{k=1}^K n_k (\tilde{\eta}_{\theta,k})^i, \quad i = 2, -5, -9.$$

By similar argument as S.4, the online estimates $\tilde{\theta}^{(K)}$ satisfies

$$\frac{\tilde{\theta}^{(K)} - \theta}{\tilde{\theta}^{(K)} - \theta} = O_p\left(1 + \frac{1}{J} + \frac{1}{J^2}\right),$$

i.e., $\tilde{\theta}^{(K)} - \theta = O_p(N_K^{-2/(d+6)})$, where J is the length of the candidate bandwidth sequence for estimating θ . Substituting the expression of $\tilde{h}^{(K)}$ and $h_*^{(K)}$ in (S.4), we obtain

$$\frac{\tilde{h}^{(K)} - h_*^{(K)}}{h_*^{(K)}} = \frac{1}{d+4} \left(\frac{\theta}{\tilde{\theta}^{(K)}}\right)^{\frac{1}{d+4}} \left(\frac{\tilde{\nu}^{(K)} - \nu}{\nu}\right) - \frac{1}{d+4} \frac{\tilde{\theta}^{(K)} - \theta}{\theta}.$$

When $h_\nu^{(K)} \asymp N_K^{-1/5}$, $\tilde{\sigma}^2(x) - \sigma^2(x) = O_p(N_K^{-2/5})$ and then $\tilde{\nu} - \nu = O_p(N_K^{-2/5})$, and hence the convergence rate of $\tilde{h}^{(K)}$ is dominated by $\tilde{\theta}^{(K)} - \theta$. \blacksquare

The online estimate $\tilde{h}^{(K)}$ attains the same convergence rate as the batch estimate $\hat{h}^{(K)}$. Similarly, in functional data analysis, one can prove that $\tilde{\theta}_\mu^{(K)} - \theta_\mu = O_p(S_{K,1}^{-2/7})$ and $\tilde{\theta}_\gamma^{(K)} - \theta_\gamma = O_p(S_{K,2}^{-1/4})$, and the convergence rate of $\tilde{h}_\mu^{(K)}$ and $\tilde{h}_\gamma^{(K)}$ are dominated by $\tilde{\theta}_\mu^{(K)} - \theta_\mu$ and $\tilde{\theta}_\gamma^{(K)} - \theta_\gamma$ respectively, in the four cases listed in Appendix B. Specifically, when estimating μ based on sparse data, from Fan and Yao (1998), when $h_\nu^{(K)} = O_p(S_{K,1}^{-1/5})$, $\tilde{r}^{(K)}(t) - r(t) = O_p(S_{K,1}^{-2/5})$

and then $\tilde{\nu}_\mu^{(K)} - \nu_\mu = O_p(S_{K,1}^{-2/5})$. For γ , one can prove that the convergence rate of $\tilde{\nu}_\gamma^{(K)}$ is dominated by $\check{E}_K \Phi(s)\Phi(t)$ in step 3 of case (b) in Appendix B which equals $O_p(S_{K,2}^{-1/3})$. For dense data, we need only check the order of $\tilde{\sigma}_K^2$. From Zhang and Chen (2007), $\check{\varepsilon}_{kij} - \varepsilon_{kij}$ in step 2 of (c) in Appendix B is of order $(S_{K,1})^{-2/25}$, recall that $\hat{\varepsilon}_{kij}^2$ is a debiased version of $\check{\varepsilon}_{kij}$ and $\tilde{\sigma}_K^2$ is the average of $\hat{\varepsilon}_{kij}^2$, then $\tilde{\sigma}_K^2 - \sigma^2 = O_p(S_{K,1}^{-12/25})$. In all the cases, estimates of θ are dominated by those of ν , which completes the proof of Theorem 3.

S.4. PROOF OF THEOREM 4

Note that the asymptotic distributions of $\tilde{\mu}$ and $\tilde{\gamma}$ have the consistent form with the standard d -dimensional local linear regression, i.e., the bias is of $O_p(h^2)$ and the variance is of $O_p(N^{-1}h^{-d})$, where $d = 1$ for mean estimation and $d = 2$ for covariance estimation. The lower bound of the relative efficiency also holds for the standard d -dimensional local linear regression. We adopt the notations of the d -dimensional local linear in the proof for notation conciseness.

Proof. Recall that $\tilde{h}^{(K)} = \eta_1^{(K)} > \dots > \eta_L^{(K)}$, without loss of generality, we assume $\eta_l^{(K)} = g(l)\tilde{h}^{(K)}$ where $g(l) \leq g(1) = 1$, $l = 1, \dots, L$. Note that $\tilde{h}^{(K)} - \hat{h}^{(K)} = o_p(N_K^{-1/(d+4)})$ and the candidate bandwidth sequence satisfies

$$\frac{\tilde{\eta}_k^{(K)}}{\tilde{h}^{(K)}} = g(l) \left(\frac{N_k}{N_K} \right)^{-\frac{1}{d+2}} + o_p(1).$$

Write $g(l) = \{g(l)^{1/\lambda}\}^\lambda$ and note that N_k/N_K grows linearly with respect to k , then λ shall be $1/(d+2)$ and the optimal $g(l)^{1/\lambda}$ shall be linear between $(0, 1)$. Hence, the optimal η_l is as in (19).

Next we derive the asymptotic lower bound for the relative efficiency of our local linear estimator. Writing $\hat{h}^{(K)} = h_*^{(K)} \{1 + O_p(N_K^{-2/(d+6)})\}$, one can derive

$$eff(\tilde{m})^{-1} = \frac{d}{d+4} \left\{ \sum_{k=1}^K \frac{n_k}{N_K} \left(\frac{\tilde{\eta}_k^{(K)}}{\tilde{h}^{(K)}} \right)^2 \right\}^2 + \frac{4}{d+4} \left\{ \sum_{k=1}^K \frac{n_k}{N_K} \left(\frac{\tilde{\eta}_k^{(K)}}{\tilde{h}^{(K)}} \right)^{-1} \right\} + O_p \left(N_K^{-\frac{2}{6+d}} \right). \quad (\text{S.5})$$

As shown in Figure 1, when K tends large, there exists a breakpoint K_0 : when $k \leq K_0$, $\tilde{\eta}_k$ equals the last candidate $(1/L)^{1/(d+4)}\tilde{h}_k$ for the limit number of candidates and for $k > K_0$,

there are sufficient and close candidates to make a choice. Hence

$$\frac{\tilde{\eta}_k^{(K)}}{\tilde{h}^{(K)}} = \begin{cases} L^{-\frac{1}{d+4}} \cdot \left(\frac{N_k}{N_K}\right)^{-\frac{1}{d+4}} + O_p\left(N_k^{-\frac{2}{6+d}}\right), & k \leq K_0 \\ \left\{1 + \left(\frac{N_k}{N_K} - \frac{l_k}{L}\right)\right\}^{\frac{1}{d+4}} + O_p\left(N_k^{-\frac{2}{6+d}}\right), & k > K_0, \end{cases}$$

where $l_k = \operatorname{argmin}_{1 \leq l \leq L} |N_k/N_K - l/L|$. Then when $k \leq K_0$,

$$\begin{aligned} \sum_{k=1}^{K_0} \frac{n_k}{N_K} \left(\frac{\tilde{\eta}_k^{(K)}}{\tilde{h}^{(K)}}\right)^2 &= \frac{d+4}{d+2} \cdot L^{-\frac{2}{d+4}} \left(\frac{N_{K_0}}{N_K}\right)^{1-\frac{2}{d+4}} + O_p\left(N_k^{-\frac{2}{6+d}}\right), \\ \sum_{k=1}^{K_0} \frac{n_k}{N_K} \left(\frac{\tilde{\eta}_k^{(K)}}{\tilde{h}^{(K)}}\right)^{-d} &= \frac{d+4}{2d+4} \cdot L^{\frac{d}{d+4}} \left(\frac{N_{K_0}}{N_K}\right)^{1+\frac{d}{d+4}} + O_p\left(N_k^{-\frac{2}{6+d}}\right). \end{aligned}$$

and when $k > K_0$,

$$\begin{aligned} \sum_{k=K_0+1}^K \frac{n_k}{N_K} \left(\frac{\tilde{\eta}_k^{(K)}}{\tilde{h}^{(K)}}\right)^2 &= 1 - \frac{N_{K_0}}{N_K} + O_p\left(N_K^{-\frac{2}{6+d}}\right), \\ \sum_{k=K_0+1}^K \frac{n_k}{N_K} \left(\frac{\tilde{\eta}_k^{(K)}}{\tilde{h}^{(K)}}\right)^{-1} &= 1 - \frac{N_{K_0}}{N_K} + O_p\left(N_K^{-\frac{2}{6+d}}\right). \end{aligned}$$

The property of breakpoint K_0 also guarantees that

$$\left(\frac{1}{L}\right)^{\frac{1}{d+4}} \sum_{k=1}^{K_0} \frac{n_k}{N_{K_0}} \tilde{h}^{(k)} \geq \tilde{h}^{(K)},$$

which is equivalent to

$$\left(\frac{1}{L}\right)^{\frac{1}{d+4}} \sum_{k=1}^{K_0} \frac{n_k}{N_{K_0}} \frac{\tilde{h}^{(k)}/h_*^{(k)}}{\tilde{h}^{(K)}/h_*^{(K)}} \frac{h_*^{(k)}}{h_*^{(K)}} = \left(\frac{1}{L}\right)^{\frac{1}{d+4}} \sum_{k=1}^{K_0} \frac{n_k}{N_{K_0}} \{1 + O_p(N_K^{-\frac{2}{6+d}})\} \left(\frac{N_k}{N_K}\right)^{\frac{1}{d+4}} \geq 1,$$

and under Assumption (A.5), we obtain

$$\frac{N_{K_0}}{N_K} \leq \left(\frac{d+4}{d+3}\right)^{d+4} \frac{1}{L} + O_p\left(N_K^{-\frac{2}{6+d}}\right).$$

Denote $\rho = N_{K_0}/N_K$, one can derive from (S.5) that

$$\begin{aligned} \operatorname{eff}(\tilde{m})^{-1} &= \frac{d}{d+4} \left\{ \frac{d+4}{d+2} \cdot L^{-\frac{2}{d+4}} \rho^{1-\frac{2}{d+4}} + 1 - \rho \right\}^2 \\ &\quad + \frac{4}{d+4} \left\{ \frac{d+4}{2d+4} \cdot L^{\frac{d}{d+4}} \rho^{1+\frac{d}{d+4}} + 1 - \rho \right\} + O_p\left(N_K^{-\frac{2}{6+d}}\right). \end{aligned} \quad (\text{S.6})$$

Note that (S.6) is strictly increasing with respect to $\rho \in \left[0, \{(d+4)/(d+3)\}^{d+4}/L\right]$. Hence we have

$$eff(\tilde{m})^{-1} \leq 1 + \frac{2c_1d + 4c_2}{d+4}L^{-1} + \frac{d}{d+4}c_1^2L^{-2},$$

where

$$c_1 = \frac{(d+4)^{d+3}}{(d+3)^{d+2}(d+2)} - \left(\frac{d+4}{d+3}\right)^{d+4}, c_2 = \frac{(d+4)^{2d+5}}{(d+3)^{2d+4}(2d+4)} - \left(\frac{d+4}{d+3}\right)^{d+4}.$$

This completes the proof of Theorem 4. ■

REFERENCES

- Fan, J. and Yao, Q. (1998), ‘Efficient estimation of conditional variance functions in stochastic regression’, *Biometrika* **85**(3), 645–660.
- Wand, M. P. and Jones, M. C. (1994), *Kernel Smoothing*, Chapman and Hall.
- Zhang, J. and Chen, J. (2007), ‘Statistical inferences for functional data’, *The Annals of Statistics* **35**(3), 1052–1079.
- Zhang, X. and Wang, J.-L. (2016), ‘From sparse to dense functional data and beyond’, *The Annals of Statistics* **44**(5), 2281–2321.